

**Solution 1** Below we use the fact that the margins of the copula are uniform, and if  $U \sim U(0, 1)$ , then  $E(U) = 1/2$  and  $\text{var}(U) = 1/12$ .

(a) By the i.i.d. assumption, using that  $P(U_1 \geq V_1) = P(U_1 \leq V_1) = 1/2$  we compute

$$\begin{aligned} \tau_{\text{corr}}(I\{U_1 > V_1\}, I\{U_2 > V_2\}) &= \frac{E(I\{U_1 > V_1\}I\{U_2 > V_2\}) - E(I\{U_1 > V_1\})E(I\{U_2 > V_2\})}{(E(I\{U_1 > V_1\}) - E(I\{U_1 > V_1\})^2)^{1/2}(E(I\{U_2 > V_2\}) - E(I\{U_2 > V_2\})^2)^{1/2}} \\ &= \frac{E(I\{U_1 > V_1\}I\{U_2 > V_2\}) - (1/2)^2}{1/4} = 4E(I\{U_1 > V_1\}I\{U_2 > V_2\}) - 1, \end{aligned}$$

where  $E(I\{U_1 > V_1\}I\{U_2 > V_2\}) = E(E(I\{U_1 > V_1\}I\{U_2 > V_2\} | (U_1, U_2))) = E(C(U_1, U_2))$ .

(b)

$$\rho = \text{corr}(U_1, U_2) = \frac{E(U_1 U_2) - E(U_1)E(U_2)}{\text{var}(U_1)^{1/2}\text{var}(U_2)^{1/2}} = \frac{E(U_1 U_2) - \frac{1}{4}}{\frac{1}{12}} = 12E(U_1 U_2) - 3.$$

(c) When  $U_1$  and  $U_2$  are independent, we get a zero in the numerator in the second line of (a).

Similarly in the first line in (b),  $E(U_1 U_2) - 1/4 = E(U_1)E(U_2) - 1/4 = 0$ .

### Solution 2

(a) The copula is defined using  $F_d^{-1}(u_d) = -1/\log u_d$  as

$$C(u_1, \dots, u_D) = F\{F_1^{-1}(u_1), \dots, F_1^{-1}(u_D)\} = \exp\{-V(-1/\log u_1, \dots, -1/\log u_D)\}, \quad 0 < u_1, \dots, < u_D.$$

(b) In terms of  $F$  we have  $F^t(tz) = F(z)$ , which when written as  $F^t(z) = F(z/t)$  gives  $C^t(u) = F^t(-1/\log u) = F(-1/t \log u) = F(-1/\log u^t) = C(u^t)$ . Replacing  $u$  by  $u^{1/t}$  gives the equation  $C^t(u^{1/t}) = C(u)$ , as required.

(c) It is easily checked that  $C_1$  and  $C_2$  are max-stable, but  $C_3$  is not.

### Solution 3

(a) Since the marginal distributions are unit Fréchet, we must have

$$G(z, \infty, \dots, \infty) = \exp\{-V(z, \infty, \dots, \infty)\} = \exp(-1/z), \quad z > 0,$$

i.e.,  $V(z, \infty, \dots, \infty) = 1/z$ , and this holds for any permutation of the arguments. For the max-stability we have

$$P(Z_d \leq b_t + a_t z)^t = [\exp\{-1/(b + t + a_t z)\}]^t = \exp(-1/z)$$

if we set  $b_t = 0$  and  $a_t = t$ , and this obviously holds for every  $z, t > 0$ .

(b) In the multivariate case we have  $a_t = (a_t^1, \dots, a_t^D)$  and  $b_t = (b_t^1, \dots, b_t^D)$ , say. If

$$G(b_t^1 + a_t^1 z_1, \dots, b_t^D + a_t^D z_D)^t = G(z_1, \dots, z_D), \quad (z_1, \dots, z_D) \in \mathcal{E}^*, t > 0,$$

then we see by replacing all but one of the  $z$ s by  $\infty$  that we must have  $a_t^d = t$  and  $b_t^d = 0$  for each  $d$ ; the marginal distributions must be max-stable with the same choice of  $a$ s and  $b$ s. Therefore

$$\begin{aligned} G(tz_1, \dots, tz_D)^t &= \exp\{-tV(tz_1, \dots, tz_D)\} \\ &= G(z_1, \dots, z_D) = \exp\{-V(z_1, \dots, z_D)\}, \quad (z_1, \dots, z_D) \in \mathcal{E}^*, t > 0, \end{aligned}$$

which gives the result.

For the second part, note that

$$\begin{aligned}
P\{\max(Z_1, \dots, Z_D) \leq z\} &= P(Z_1 \leq z, \dots, Z_D \leq z) \\
&= \exp\{-V(z, \dots, z)\} \\
&= \exp\{-zV(z, \dots, z)/z\} \\
&= \exp\{-V(1, \dots, 1)/z\} \\
&= \exp(-\theta_{\mathcal{D}}/z), \quad z > 0.
\end{aligned}$$

(c) Note that when  $z \rightarrow \infty$ ,  $P(Z_1 > z) = 1 - \exp(-1/z) = 1/z + O(1/z^2)$  and

$$\begin{aligned}
P(Z_1 > z, Z_2 > z) &= 1 - P(Z_1 \leq z) - P(Z_2 \leq z) + P(Z_1 \leq z, Z_2 \leq z) \\
&= 1 - 2\exp(-1/z) + \exp\{-V(z, z)\} \\
&= 1 - 2(1 - 1/z) + \{1 - V(1, 1)/z\} + O(1/z^2) \\
&= (2 - \theta)/z + O(1/z^2),
\end{aligned}$$

so

$$\chi = \lim_{z \rightarrow \infty} \frac{P(Z_1 > z, Z_2 > z)}{P(Z_1 > z)} = \lim_{z \rightarrow \infty} \frac{(2 - \theta)/z + O(1/z^2)}{1/z + O(1/z^2)} = 2 - \theta.$$

In the two cases stated we have (independent variables),  $\theta = V(1, 1) = 1/1 + 1/1 = 2$ , so  $\chi = 0$ , and (totally dependent variables)  $\theta = V(1, 1) = 1/\min(1, 1) = 1$ , so  $\chi = 1$ . Thus  $\theta$  can be seen as a measure of the asymptotic dependence between the two variables, with smaller values for higher dependence, and  $\theta = D$  corresponding to independence.

**Solution 4** Setting  $z_1 = \infty$  shows that the margin for  $z_2$  is unit Fréchet, and the result for  $z_1$  follows by symmetry. The density function can be written in the form  $(z_1 z_2)^{-2} V(z_1, z_2) \exp\{-V(z_1, z_2)\}$ , which is positive, so this is a joint density. As  $t \rightarrow \infty$ ,

$$F(tz_1, tz_2)^t = \exp\left[-\left\{z_1^{-1} + z_2^{-1} + (tz_1 z_2)^{-1}\right\}\right] \rightarrow \exp\left\{-\left(z_1^{-1} + z_2^{-1}\right)\right\},$$

so it is not max-stable, and the limiting distribution for rescaled maxima is independence. For it to correspond to a Poisson process intensity,  $V$  should correspond to a measure on  $\mathcal{E}^*$ , so the measure of any rectangle  $(z'_1, z_1) \times (z'_2, z_2)$  must be positive, i.e.,

$$V(z_1, z'_2) + V(z'_1, z_2) - V(z_1, z_2) - V(z'_1, z'_2) > 0, \quad 0 < z'_1 < z_1, 0 < z'_2 < z_2.$$

However this equals  $-(1/z'_1 - 1/z_1)(1/z'_2 - 1/z_2)$ , which is negative, so there is no such Poisson process.