

# MATH562 – Fall 2025

## Problem Set: Week 5

- \*1. A location-scale model is of the form  $y_0 = \eta + \tau v_0$ , where  $v_0$  has a known distribution, with location and scale parameters  $\eta \in \mathbb{R}$  and  $\tau > 0$ . A location estimator  $m(y)$  and scale estimator  $s(y)$ , based on a random sample  $y = (y_1, \dots, y_n)$  with the  $y_i$  iid. distributed as  $y_0$ , are said to be *equivariant* if they satisfy

$$\begin{aligned} m(a + b y) &= m(a + b y_1, \dots, a + b y_n) = a + b m(y_1, \dots, y_n) = a + b m(y), \\ s(a + b y) &= s(a + b y_1, \dots, a + b y_n) = b s(y_1, \dots, y_n) = b s(y), \quad a \in \mathbb{R}, b > 0. \end{aligned}$$

- (a) Show that  $q_1(y, \eta) = \{m(y) - \eta\}/s(y)$  and  $q_2(y, \tau) = s(y)/\tau$  are pivots, and explain how to use them to construct confidence intervals for  $\eta$  and  $\tau$ .
- (b) Deduce that pivots can be formed by taking
- (i)  $m_1(y) = \bar{y}$  and  $s_1(y) = \left\{ \sum_j (y_j - \bar{y})^2 \right\}^{1/2}$ ,
  - (ii)  $m_2(y) = \text{median}(y)$  and  $s_2(y) = \text{IQR}(y)$ , the interquartile range,
- whatever the distribution of  $v_0$ . Briefly discuss the corresponding confidence intervals.
- (c) Provide a pivot that could be used to construct a prediction interval for  $y_0$  based on  $y_1, \dots, y_n$ .
2. Find an  $\alpha$ -level confidence interval for  $Y_+$ , when  $Y_1, \dots, Y_n, Y_+ \stackrel{i.i.d.}{\sim} \exp(\lambda)$ . How does the interval change as  $n \rightarrow \infty$ . *Hint: if  $E \sim \exp(1)$  then  $2E \sim \chi_2^2$ .*
3. (a) An experiment consists of observing the number of success  $y_1$  in a fixed number  $n_1$  of independent Bernoulli trials with unknown success probability  $\theta \in (0, 1)$ . Show that the corresponding density is

$$f(y_1 | \theta) = \binom{n_1}{y_1} \theta^{y_1} (1 - \theta)^{n_1 - y_1}, \quad y_1 \in \mathcal{S}_\infty = \{0, 1, \dots, n_1\}.$$

If prior information on  $\theta$  can be summarised by the *beta density*

$$f(\theta) = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1}, \quad 0 < \theta < 1, \quad a, b > 0,$$

show that the posterior density for  $\theta$  given  $y_1$  is

$$f(\theta | y_1) = \frac{\Gamma(n_1 + a + b)}{\Gamma(y_1 + a)\Gamma(n_1 - y_1 + b)} \theta^{y_1 + a - 1} (1 - \theta)^{n_1 - y_1 + b - 1}, \quad 0 < \theta < 1.$$

- (b) Another experiment conducts independent Bernoulli trials until there are  $y_2$  successes, at which point there have been  $n_2$  trials. Show that the corresponding density is

$$f(n_2 | \theta) = \binom{n_2 - 1}{y_2 - 1} \theta^{y_2} (1 - \theta)^{n_2 - y_2}, \quad n_2 \in \mathcal{S}_\infty = \{y_2, y_2 + 1, \dots\}.$$

Without doing any calculations, write down the posterior density for  $\theta$  based on the prior in (a).

- (c) Show that if  $y_1 = y_2$  and  $n_1 = n_2$ , then Bayesian inferences based on either of the two experiments will be identical, i.e., they do not take into account the different reference sets  $\mathcal{S}_\infty$  and  $\mathcal{S}_\infty$ .

- (d) Consider testing the hypothesis that  $\theta = \frac{1}{2}$  against the alternative that  $\theta < \frac{1}{2}$ . Explain why the respective significance levels for the experiments in (a) and (b) would be

$$\sum_{y=0}^{y_1} \binom{n_1}{y} 2^{-n_1}, \quad \sum_{n=n_2}^{\infty} \binom{n-1}{y_2-1} 2^{-n},$$

and evaluate these when  $n_1 = n_2 = 12$ ,  $y_1 = y_2 = 3$ . How does this compare with (c)?

*Hint:* Recall that for  $\alpha > 0$  the gamma function is defined as  $\Gamma(\alpha) = \int_0^{\infty} u^{\alpha-1} e^{-u} du$ , and that  $\Gamma(\alpha+1) = \alpha\Gamma(\alpha)$ ,  $\Gamma(n+1) = n!$  for  $n \in \{1, 2, \dots\}$ , and  $\Gamma(1/2) = \sqrt{\pi}$ .

4. Consider the shoe data example from the Slides (Part 1).

- (a) Show that the average  $\bar{D}$  has mean  $\theta$  and variance  $\sigma^2 = m^{-2} \sum_{j=1}^m c_j^2$ .
- (b) Show that  $S^2$  has mean  $\sigma^2$  and hence can be used to estimate  $\sigma^2$ .
- (c) Extend this discussion to a balanced design in which  $m$  is even,  $I_j = \pm 1$  and  $\sum_{j=1}^m I_j = 0$  but the allocation is otherwise completely at random. This ensures that materials A and B appear equally often on the left and right shoes.