

Solution 1. (a) (1 point) The sample space for this experiment can be written

$$\Omega = \{(0, 1), (0, 2), (0, 3), (0, 4), (0, 5), (0, 6), (1, 0)\},$$

where the ordered pairs represent (score from coin, score from die).

(b) (2 points) Let S be the random variate corresponding to the total score of the coin and the die. The probability that S is odd is

$$\begin{aligned} \Pr('S \text{ is odd}') &= \Pr(\{S = 1\} \cup \{S = 3\} \cup \{S = 5\}) \\ &= \Pr(S = 1) + \Pr(S = 3) + \Pr(S = 5) \\ &= \Pr\{(0, 1)\} + \Pr\{(1, 0)\} + \Pr\{(0, 3)\} + \Pr\{(0, 5)\} \\ &= \frac{1}{2} \times \frac{1}{6} + \frac{1}{2} + \frac{1}{2} \times \frac{1}{6} + \frac{1}{2} \times \frac{1}{6} = \frac{3}{4}. \end{aligned}$$

(c) (2 points) The probability that S equals 3 or less is

$$\begin{aligned} \Pr(S \leq 3) &= \Pr(\{S = 1\} \cup \{S = 2\} \cup \{S = 3\}) \\ &= \Pr(S = 1) + \Pr(S = 2) + \Pr(S = 3) \\ &= \Pr\{(0, 1)\} + \Pr\{(1, 0)\} + \Pr\{(0, 2)\} + \Pr\{(0, 3)\} \\ &= \frac{1}{2} \times \frac{1}{6} + \frac{1}{2} + \frac{1}{2} \times \frac{1}{6} + \frac{1}{2} \times \frac{1}{6} = \frac{3}{4}. \end{aligned}$$

(d) (2 points) Denote H the event that the coin showed heads. The probability that the coin showed heads if the total score equals 3 or less is

$$\Pr(H | S \leq 3) = \frac{\Pr(H \cap S \leq 3)}{\Pr(S \leq 3)} = \frac{\Pr\{(1, 0)\}}{\Pr(S \leq 3)} = \frac{1/2}{3/4} = 2/3.$$

Solution 2. (a) (3 points) The probability of winning the lottery on the next draw is

$$\Pr('win') = \frac{\binom{5}{5} \binom{2}{2}}{\binom{50}{5} \binom{11}{2}} = 1/116'531'800 \approx 8.581 \times 10^{-09}.$$

(b) (1 points) Denote X the number of winners at the next draw. Since each person has one chance in 116'531'800 to win and the people play independently, $X \sim B(n = 80'000'000, p = 1/116'531'800)$, and hence the expected number of winners at the next draw is

$$E(X) = np = \frac{80'000'000}{116'531'800} = 0.6865.$$

(c) (3 points) Here the law of small numbers applies (n is very large and p very small), so X follows approximately a Poisson distribution with parameter $\lambda = np = 0.6865$. So,

$$\begin{aligned} \Pr(X = 0) &= \frac{0.6865^0}{0!} e^{-0.6865} \approx 0.5033, \\ \Pr(X = 1) &= \frac{0.6865^1}{1!} e^{-0.6865} \approx 0.3455, \end{aligned}$$

and $\Pr(X \geq 2) = 1 - \Pr(X = 0) - \Pr(X = 1) \approx 0.1512$.

Alternatively, from the binomial distribution we have

$$\begin{aligned} \Pr(X = 0) &= \binom{80'000'000}{0} \left(\frac{1}{116'531'800}\right)^0 \left(1 - \frac{1}{116'531'800}\right)^{80'000'000} \approx 0.5033, \\ \Pr(X = 1) &= \binom{80'000'000}{1} \left(\frac{1}{116'531'800}\right)^1 \left(1 - \frac{1}{116'531'800}\right)^{79'999'999} \approx 0.3455, \end{aligned}$$

and $\Pr(X \geq 2) = 1 - \Pr(X = 0) - \Pr(X = 1) \approx 0.1512$.

- (d) **(2 points)** Let N denote the number of draws until the next winner. The event $\{N = n\}$ happens if there is no winner at the $n - 1$ first draws and there is at least one winner at the n th draw. Each draw is a Bernoulli trial with probability of ‘success’ $\tilde{p} = \Pr(X \geq 1) = 0.4967$. Therefore, the number of draws N until the first win follows a geometric distribution of parameter \tilde{p} , and mass function

$$\Pr(N = n) = \tilde{p}(1 - \tilde{p})^{n-1}, \quad n = 1, 2, \dots$$

- Solution 3. (a) (4 points)** Let R and S denote respectively the run times of the first and second jobs. Since they are uniformly distributed from zero to three hours, their cumulative distribution functions are

$$F_R(z) = F_S(z) = \begin{cases} 0, & z \leq 0, \\ \frac{z}{3}, & 0 < z < 3, \\ 1, & z \geq 3. \end{cases}$$

The cumulative distribution function of the longest run time $X = \max(R, S)$ is

$$\begin{aligned} F_X(x) &= \Pr(X \leq x) = \Pr\{\max(R, S) \leq x\} = \Pr(R \leq x, S \leq x) \\ &= \Pr(R \leq x)\Pr(S \leq x) = F_R(x)F_S(x). \end{aligned}$$

So,

$$F_X(x) = \begin{cases} 0, & z \leq 0, \\ \frac{z^2}{9}, & 0 < z < 3, \\ 1, & z \geq 3. \end{cases}$$

Therefore, the probability density function of X is

$$f_X(x) = \frac{d}{dx}F_X(x) = \begin{cases} 2x/9, & 0 < x < 3, \\ 0, & \text{elsewhere.} \end{cases}$$

- (b) **(2 points)** The expected waiting time until both processors are free is

$$E(X) = \int_0^3 x \frac{2x}{9} dx = \frac{2}{9} \frac{x^3}{3} \Big|_0^3 = 2 \text{ hours.}$$

- (c) **(2 points)** The probability that the longest run time was less than one hour given that both jobs were done after two hours is

$$\Pr(X < 1 \mid X < 2) = \frac{\Pr(X < 1, X < 2)}{\Pr(X < 2)} = \frac{\Pr(X < 1)}{\Pr(X < 2)} = \frac{1/9}{4/9} = 1/4.$$

- (d) **(2 points)** The expected running time until both cores are free if the computer is still running after two hours is

$$\begin{aligned} E(X \mid X > 2) &= \frac{E\{XI(X > 2)\}}{\Pr(X > 2)} \\ &= \frac{1}{1 - \Pr(X \leq 2)} \int_0^3 xI(x > 2) \frac{2x}{9} dx \\ &= \frac{9}{5} \int_2^3 x \frac{2x}{9} dx \\ &= \frac{2}{5} \frac{x^3}{3} \Big|_2^3 = \frac{38}{15} \approx 2.53. \end{aligned}$$

So, the expected remaining waiting time until both cores are free if the computer is still running after two hours is $E(X - 2 \mid X > 2) = 38/15 - 2 = 8/15 \approx 0.53$ hour.

Solution 4. (a) (4 points) Let T denote the download time in hours. The probability that the download time exceeds t hours is

$$\Pr(T > t) = \int_t^\infty 10 \exp(-10x) dx = -\exp(-10x)|_t^\infty = \exp(-10t).$$

The conditional probability that the download time exceeds $t + x$ hours, given that it exceeds t hours, is

$$\begin{aligned} \Pr(T > t + x | T > t) &= \frac{\Pr(T > t + x, T > t)}{\Pr(T > t)} = \frac{\Pr(T > t + x)}{\Pr(T > t)} \\ &= \frac{\exp(-10t - 10x)}{\exp(-10t)} = \exp(-10x), \quad t, x > 0. \end{aligned}$$

Note that $\Pr(T > t + x | T > t) = \Pr(T > x)$. So the download time T exhibits the lack of memory property, which one should expect since T follows an exponential distribution with parameter $\lambda = 10$.

(b) (4 points) First, we find the constant c . To be a valid density, the joint density must integrate to one, so

$$\begin{aligned} 1 &= \int_0^\infty \int_0^\infty c \exp(-10x - 20y) dx dy = c \int_0^\infty \exp(-10x) dx \int_0^\infty \exp(-20y) dy \\ &= c \left[-\frac{1}{10} \exp(-10x) \right]_0^\infty \left[-\frac{1}{20} \exp(-20y) \right]_0^\infty \\ &= c \frac{1}{200}, \end{aligned}$$

giving $c = 200$.

The marginal density of Y is

$$\begin{aligned} f_Y(y) &= \int_0^\infty f_{X,Y}(x, y) dx = 200 \exp(-20y) \int_0^\infty \exp(-10x) dx \\ &= 200 \exp(-20y) \left[-\frac{1}{10} \exp(-10x) \right]_0^\infty = 20 \exp(-20y), \quad y > 0. \end{aligned}$$

The random variates X and Y are independent since their support is a Cartesian product, $(0, \infty) \times (0, \infty)$, and in the support, the joint density factorises :

$$f_{X,Y}(x, y) = 200 \exp(-10x - 20y) = 10 \exp(-10x) \times 20 \exp(-20y) = f_X(x) \times f_Y(y);$$

clearly $X \sim \exp(\lambda = 10)$ and $Y \sim \exp(\lambda = 20)$.

(c) (4 points) The probability that $Y < X$ is

$$\begin{aligned} \Pr(Y < X) &= \int_{y=0}^\infty dy \int_{x=y}^\infty dx f_{X,Y}(x, y) \\ &= \int_{y=0}^\infty dy 20 \exp(-20y) \int_{x=y}^\infty dx 10 \exp(-10x) \\ &= \int_{d=0}^\infty dy 20 \exp(-20y) [-\exp(-10x)]_{x=y}^\infty \\ &= 20 \int_0^\infty \exp(-30y) dy = 20 \left[-\frac{1}{30} \exp(-30y) \right]_0^\infty = \frac{2}{3}. \end{aligned}$$