

1

RECAP: THE MEAN OF MEANS, THE VARIANCE OF THE MEANS.

$$Z_1 = \bar{X}_{S1} = 1200 \text{ kg} / Z_2 = \bar{X}_{S2} = 1310 \text{ kg}$$

$$N_{S1} = 50, N_{S2} = 50, \text{ etc...}$$

$$N_{S1} = 100, N_{S2} = 100, N_{S3} = 100, \text{ etc...}$$

$$Y_1 = \bar{X}_{S1} = 1300 \text{ kg}, Y_2 = \bar{X}_{S2} = 1310 \text{ kg}, Y_3 = \bar{X}_{S3} = 1360 \text{ kg}$$

$$\tilde{S}_1^2 = (200 \text{ kg})^2, \tilde{S}_2^2 = (220 \text{ kg})^2, \dots$$

$$S_1^2 \approx (200 \text{ kg})^2, S_2^2 \approx (220 \text{ kg})^2$$

POPULATION:
"ALL CARS IN ITALY"

$$N_{\text{POP}} = 40 \text{ MILLION}$$

TRUE MEAN OF THE WEIGHT:

$$\mu = 1330 \text{ kg}$$

$$\text{VARIANCE: } \sigma^2 = (200 \text{ kg})^2$$

Every sample containing 100 points/cars has a mean $\bar{X}_{Si}$ which we call Y_i . The collection of all these means (there would be 400'000 means) has $E(Y) = \mu$ (naturally the mean of all means gives the population mean)

and $\text{Var}(Y) = \frac{\sigma^2}{N_s} = \frac{\sigma^2}{100}$ i.e. the variance of the means is 100x less than the population variance.
(see Central limit theorem, CLT.)

The CLT also tells us that Y has an approximately Gaussian distribution.

What if we had chosen $N_{S1} = 50$. Let's call these means Z_i .

Now $E(Z_i) = \mu$ $\text{Var}(Z_i) = \frac{\sigma^2}{N_{S1}} = \frac{\sigma^2}{50}$

What if we combine two samples, say Z_i and Z_j (which would correspond to Y_1), and average those?

$$E\left(\frac{Z_i + Z_j}{2}\right) = \frac{1}{2}(E(Z_i) + E(Z_j)) = \frac{1}{2}(\mu + \mu) = \mu = E(Y_i) \rightarrow \text{Same.}$$

$$\text{Var}\left(\frac{Z_i + Z_j}{2}\right) = \frac{1}{4} \text{Var}(Z_i + Z_j) = \frac{1}{4}(\text{Var}(Z_i) + \text{Var}(Z_j)) = \frac{1}{4}\left(\frac{\sigma^2}{50} + \frac{\sigma^2}{50}\right) = \frac{\sigma^2}{100} = \text{Var}(Y_i) \rightarrow \text{Same}$$

Var is quadratic! → Z_i and Z_j are independent

[2] We measured a sample, and got $\bar{x}_S$ and s^2 . What now?

- Our best guess of μ is $\bar{x}_S = \frac{1}{N_S} \sum_{i=1}^{N_S} x_i$
- Our best guess of σ^2 is s^2 , where s^2 is the unbiased sample variance and $s^2 = \frac{N_S}{N_S - 1} \tilde{s}^2$ and $\tilde{s}^2 = \frac{1}{N_S} \sum_{i=1}^{N_S} (x_i - \bar{x}_S)^2$ is the simple sample variance because

$E(\tilde{s}^2) = \sigma^2 - \frac{\sigma^2}{N}$ i.e. it has a bias of $-\frac{\sigma^2}{N}$ so it will probably under estimate the true σ .

- Given $\bar{x}_S$ and assuming we know the true σ^2 , how confident can I be in my knowledge of the true μ ?

i.e. What is the probability $P(\bar{x}_S - a \leq \mu \leq \bar{x}_S + a)$ given $\bar{x}_S$ and N_S ?

[3] We know, from the CLT: $p(x_S) = \exp\left(-\frac{(\bar{x}_S - \mu)^2}{2(\sigma^2/N_S)}\right) \cdot \frac{1}{\sqrt{2\pi} \sigma/\sqrt{N_S}}$

Hence: $P\left(\frac{\bar{x}_S - \mu}{\sigma/\sqrt{N_S}} \leq z\right) = \Phi(z)$ (standard normal CDF)

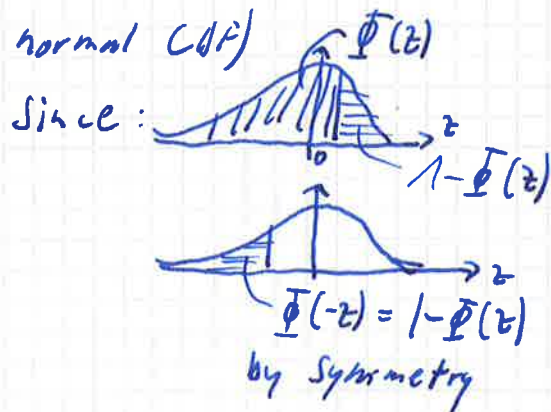
$$P\left(-z \leq \frac{\bar{x}_S - \mu}{\sigma/\sqrt{N_S}} \leq z\right) = 2\Phi(z) - 1$$

$$= P\left(-\frac{z\sigma}{\sqrt{N_S}} \leq \bar{x}_S - \mu \leq \frac{z\sigma}{\sqrt{N_S}}\right)$$

$$= P\left(-\frac{z\sigma}{\sqrt{N_S}} - \bar{x}_S \leq -\mu \leq \frac{z\sigma}{\sqrt{N_S}} - \bar{x}_S\right)$$

divide by -1 , this swaps $\leq$ to $\geq$. Then swap the left/right limit:

$$= P\left(\bar{x}_S - \frac{z\sigma}{\sqrt{N_S}} \leq \mu \leq \bar{x}_S + \frac{z\sigma}{\sqrt{N_S}}\right) = 2\Phi(z) - 1$$



So with $\bar{x}_S = 1300 \text{ kg}$, $N_S = 100$, $\sigma = 200 \text{ kg}$: " μ lies between 1270 and 1330

With a probability... $\left[130 = \bar{x}_S + \frac{z\sigma}{\sqrt{N_S}} \Rightarrow z = \frac{30}{200/\sqrt{100}} = 1.5\right] 2\Phi(1.5) - 1 = 87\%$

MJE 213 L 6 p3

[3 bis] Also we can say:

With 95% probability $[2\Phi(z) - 1 = 0.95 \Rightarrow \Phi(z) = 0.975 \Rightarrow z \approx 2$
 $\Rightarrow x_s - \frac{z\sigma}{\sqrt{n}} = 1260]$ μ will be between 1260 and 1340

$\Rightarrow$ "The 95% confidence interval is 1260 to 1340."

Hypothesis Testing: Your friend throws 10 times "6" with a die...

[4] H_A : "The die is fair" $p_1 = p_2 = \dots = p_6 = \frac{1}{6}$

$P(10 \text{ times } 6 | H_A) = \left(\frac{1}{6}\right)^{10} = 0.0000017\% \rightarrow$ seems unreasonable...

H_B : The die has $p_6 = 0.9$, $p_2 = 0.1$ others = 0

$P(10 \text{ times } 6 | H_B) = (0.9)^{10} = 35\% \rightarrow$ plausible

H_C : $P(10 \times 6 | H_C) = 0.99^{10} = 90\% \rightarrow$ also plausible...
 $\rightarrow p_6 = 0.99, p_5 = 0.01$

H_D : $p_6 = 1$ others = 0

$P(10 \times 6 | H_D) = 1^{10} = 1 \dots \rightarrow$ also plausible...?

$\hookrightarrow P(\text{outcome} | \text{HYPOTHESIS}) \neq P(\text{HYPOTHESIS TRUE})$

Hypothesis examples:
- The mean weight of cars in the USA > in Italy $\rightarrow$ YES/NO
- Smoking makes cancer more likely than not smoking $\rightarrow$ Y/N
- There is a particle at 125 GeV

[5] Hypothesis testing procedure

- 1) Get a YES/No Hypothesis, : $\left\{ \begin{array}{l} \text{"My friend cheats"} \rightarrow H_1 \\ \text{as a statement about} \\ \text{a probability distribution.} \end{array} \right. \left\{ \begin{array}{l} \text{L) "The die does not follow a Laplace} \\ \text{distribution."} \rightarrow \text{cannot calculate } P \end{array} \right.$
 $\rightarrow "H_1"$
 - 2) Get the logical opposite of H_1 , i.e. something that covers all other scenarios except H_1 . $\left\{ \begin{array}{l} \rightarrow \text{"The die does follow} \\ \text{a Laplace distribution"} \rightarrow H_0 \\ \rightarrow \text{can calculate } P \end{array} \right.$
 $\rightarrow "H_0", \text{ the "null hypothesis"}$
 - 3) Choose a level of "statistical significance" α . This is the chance you will reject H_0 (and believe H_1) ~~can~~ even if it is true. $\left\{ \begin{array}{l} \alpha = 5\% \text{ is} \\ \text{common... should} \\ \text{it be?} \end{array} \right.$
 - 4) Compute $P(\text{outcome} | H_0)$ for some measurement/observation.
 If $P < \alpha \rightarrow$ ~~can~~ reject H_0 , believe H_1
- Die example: $P(10 \times 6 | H_0) = 0.0000017\% < 5\% \rightarrow$ reject $H_0 \dots$
 $\rightarrow$ "Friend cheats"
- 2 throws only: $P(2 \times 6 | H_0) = \frac{1}{36} = 2.8\% < 5\% \rightarrow$ would you reject H_0 ?