

**Problem 1** Let  $A$  and  $B$  be  $q \times q$  matrices, and suppose that  $(A + \alpha B)^{-1}$  exists for some  $\alpha > 0$ . If  $\eta$  is an eigenvalue of  $(A + \alpha B)^{-1}A$ , show that

- (a) if  $B$  is invertible, then

$$(A + \alpha B)^{-1}A = B^{-1/2}(B^{-1/2}AB^{-1/2} + \alpha I_q)^{-1}B^{-1/2}A,$$

and deduce that  $\eta = \eta' / (\alpha + \eta')$  where  $\eta'$  is an eigenvalue of  $B^{-1/2}AB^{-1/2}$ , and

- (b) if  $A$  is invertible, then  $\eta = 1 / (1 + \alpha\eta'')$ , where  $\eta''$  is an eigenvalue of  $A^{-1/2}BA^{-1/2}$ .

**Problem 2**

- (a) If  $X$  has  $n > p$  and rank  $p$ , use its singular value decomposition to write the linear model  $y \sim (X\beta, \sigma^2 I_n)$  as  $y \sim (UD\gamma, \sigma^2 I_n)$ , and give expressions for the least squares estimators  $\hat{\beta}$  and  $\hat{\gamma}$ .
- (b) Show that the squared Euclidean distance of  $\hat{\beta}$  from  $\beta$ , i.e.,  $Q = \|\hat{\beta} - \beta\|_2^2$ , can be written as  $\|\hat{\gamma} - \gamma\|_2^2$ , where  $\hat{\gamma} = \text{diag}(1/d_1, \dots, 1/d_p, 0, \dots, 0)U^T y$ . Under what circumstances will the variance of  $\hat{\gamma}$  be large?
- (c) If  $y$  has a normal distribution, show that  $Q \stackrel{D}{=} \sum_{r=1}^p \sigma^2 Z_r^2 / d_r^2$ , where  $Z_1, \dots, Z_p \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1)$ , and hence find the mean and variance of  $Q$ .

**Problem 3**

- (a) Show directly that  $\hat{\beta}_\lambda = (X^T X + \lambda I_p)^{-1} X^T y$  minimises the ridge regression sum of squares

$$(y - X\beta)^T(y - X\beta) + \lambda\beta^T\beta,$$

when the response  $y$  and the covariate matrix  $X$  are centred, so the model does not include a vector of ones.

- (b) Use the singular value decomposition  $X = UDV^T$ , in which  $V = (v_1, \dots, v_p)$  and  $U = (u_1, \dots, u_n)$  in terms of column vectors, to show that

$$\hat{\beta}_\lambda = \sum_{d_j > 0} \frac{d_j}{d_j^2 + \lambda} u_j^T y \times v_j.$$

Give a similar expression for the fitted value  $\hat{y}_\lambda = X\hat{\beta}_\lambda$  and discuss the effect of increasing  $\lambda$  for different values of  $d_j$ .

- (c) Find expressions for  $\text{edf}_\lambda$  and for the bias and variance of  $\hat{\beta}_\lambda$  in terms of the SVD.

**Problem 4** (Centering and penalisation)

- (a) In a linear model whose response vector  $y$  has mean  $\beta_0 1_n + X\beta$ , show that

$$y - \beta_0 1_n - X\beta = y_* - (\gamma - \bar{y}) 1_n - X_*\beta,$$

with a suitable choice of  $\gamma$ , where  $1_n^\top y_* = 0$  and  $1_n^\top X_* = 0$ ; this splits  $\mathbb{R}^n$  into  $\text{span}(1_n)$  and its orthogonal subspace. Is the interpretation of  $\beta$  in this new parametrisation the same as in the original model?

- (b) Deduce that both minimisation problems

$$\min_{\beta_0, \beta} \|y - \beta_0 1_n - X\beta\|_2^2 + \lambda p(\beta), \quad \min_{\beta} \|y_* - X_*\beta\|_2^2 + \lambda p(\beta)$$

give the same estimate  $\hat{\beta}_\lambda$ , and conclude that ridge regression and lasso fits to centred response and design matrices need not include the intercept.

- (c) Show that if  $\beta_0$  is included in  $\beta$  and is penalised, then the problem is not invariant to response transformations of the form  $y \mapsto ay + b1_n$  for constants  $a > 0$  and  $b$ . Explain why invariance to such transformations is desirable.